Illinois State University

## ISU ReD: Research and eData

Faculty Publications – Mathematics

**Mathematics** 

2024

# The Law of the Iterated Logarithm for Lp-Norms of Kernel Estimators of Cumulative Distribution Functions

Fuxia Cheng Illinois State University, fcheng@ilstu.edu

Follow this and additional works at: https://ir.library.illinoisstate.edu/fpmath

Part of the Mathematics Commons

### **Recommended Citation**

Cheng, Fuxia. 2024. "The Law of the Iterated Logarithm for Lp-Norms of Kernel Estimators of Cumulative Distribution Functions" Mathematics 12, no. 7: 1063. https://doi.org/10.3390/math12071063.

This Article is brought to you for free and open access by the Mathematics at ISU ReD: Research and eData. It has been accepted for inclusion in Faculty Publications – Mathematics by an authorized administrator of ISU ReD: Research and eData. For more information, please contact ISUReD@ilstu.edu.





## Article **The Law of the Iterated Logarithm for** *L*<sub>*p*</sub>**-Norms of Kernel Estimators of Cumulative Distribution Functions**

**Fuxia Cheng** 

Department of Mathematics, Illinois State University, Normal, IL 61790, USA; fcheng@ilstu.edu; Tel.: +1-309-438-7597

**Abstract:** In this paper, we consider the strong convergence of  $L_p$ -norms ( $p \ge 1$ ) of a kernel estimator of a cumulative distribution function (CDF). Under some mild conditions, the law of the iterated logarithm (LIL) for the  $L_p$ -norms of empirical processes is extended to the kernel estimator of the CDF.

Keywords: L<sub>p</sub>-norm; LIL; kernel estimator; empirical CDF

MSC: 60F15; 62G05

#### 1. Introduction

Consider an independent identically distributed random sample  $X_1, X_2, \dots, X_n$  from a population with an unknown cumulative distribution function (CDF). For the empirical distribution function  $F_n$ , defined as follows:

$$F_n(x) = \frac{1}{n} \sum_{i=1}^n I(X_i \le x), \, \forall x \in \mathbb{R}^1,$$

with *I* denoting the indicator function, the classical Glivenko–Cantelli theorem states that  $F_n(x)$  converges almost surely (a.s.) to F(x) uniformly in  $x \in \mathbb{R}^1$ , i.e.,

$$\sup_{x\in R^1} |F_n(x) - F(x)| \to 0, \text{ a.s.}$$

The extended Glivenko–Cantelli lemma (in Fabian and Hannan 1985, pp. 80–83 [1]) provides the strong uniform convergence rate as follows:

$$\sup_{x \in R} n^{\alpha} |F_n(x) - F(x)| \to 0 \quad \text{a.s.,} \quad \text{for any } 0 < \alpha < 1/2.$$
(1)

The law of the iterated logarithm (LIL) for  $F_n(t)$ , i.e.,

$$\lim \sup_{n \to \infty} \sqrt{\frac{n}{2\log(\log n)}} \sup_{x} \left| F_n(x) - F(x) \right| = \frac{1}{2} \quad a.s.$$
(2)

was proven by Smirnov (1944) [2] and, independently, Chung (1949) [3]. Finkelstein (1971) [4] obtained the *L*<sub>2</sub>-version of the law of iterated logarithm,

$$\lim \sup_{n \to \infty} \sqrt{\frac{n}{2\log(\log n)}} \left[ \int_{-\infty}^{\infty} (F_n(x) - F(x))^2 dF(x) \right]^{1/2} = \frac{1}{\pi} \quad a.s.$$
(3)

For any  $p \ge 1$ , setting

$$C(p) = \frac{1}{2} \left(\frac{p(p+2)}{\pi}\right)^{1/2} \left(\frac{2}{p+2}\right)^{1/p} \frac{\Gamma(1/p + \frac{1}{2})}{\Gamma(1/p)},\tag{4}$$



Citation: Cheng, F. The Law of the Iterated Logarithm for *L<sub>p</sub>*-Norms of Kernel Estimators of Cumulative Distribution Functions. *Mathematics* 2024, *12*, 1063. https://doi.org/ 10.3390/math12071063

Academic Editor: Maurizio Brizzi

Received: 20 February 2024 Revised: 23 March 2024 Accepted: 28 March 2024 Published: 1 April 2024



**Copyright:** © 2024 by the author. Licensee MDPI, Basel, Switzerland. This article is an open access article distributed under the terms and conditions of the Creative Commons Attribution (CC BY) license (https:// creativecommons.org/licenses/by/ 4.0/). the law of the iterated logarithm for  $L_p$ -norm of  $F_n(x)$ ,

$$\lim \sup_{n \to \infty} \sqrt{\frac{n}{2\log(\log n)}} \left[ \int_{-\infty}^{\infty} |F_n(x) - F(x)|^p dF(x) \right]^{1/p} = C(p) \quad a.s.$$
(5)

was developped by Gajek, Kahszka, and Lenic (1996) [5]. It is easy to verify that

$$C(1) = \frac{\sqrt{3}}{6}$$
 and  $C(2) = \frac{1}{\pi}$ 

And (3) is a special case of (5) corresponding to p = 2.

Notice that there is one serious discontinuity drawback of  $F_n$ , regardless of F being continuous or discrete. To treat this deficiency of  $F_n$ , Yamato (1973) [6] proposed the following kernel distribution estimator:

$$\hat{F}(x) = \int_{-\infty}^{x} n^{-1} \sum_{i=1}^{n} k_h (u - X_i) du, \ x \in \mathbb{R},$$
(6)

in which  $h = h_n$  is the usual band width sequence of positive numbers tending to zero, k is a probability density function(PDF) called kernel, and  $k_h(u) = k(u/h)/h$ .

The aim of this paper is to provide certain conditions to guarantee the LIL of  $L_p$ -norm of  $\hat{F}$ . Some asymptotic properties of the smooth estimator  $\hat{F}$  have been established. For example, in Yamato (1973) [6], the asymptotic normality and uniform strong consistency of  $\hat{F}$  were obtained. In more general contexts, Winter (1979) [7] considered the convergence rate of perturbed empirical distribution functions. Wang, Cheng, and Yang (2013) [8] developed simultaneous confidence bands for F based on  $\hat{F}$ . The strong convergence rate of  $\hat{F}$  was considered by Cheng (2017) [9], which extended the extended Glivenko–Cantelli Lemma (1) to the kernel estimator  $\hat{F}$ .

Here, we shall continue to consider the strong convergence of a smooth estimator  $\hat{F}$  for F. More specifically, we are interested in extending the LIL of  $L_p$ -norm in (5) for  $F_n(t)$  to the kernel estimator  $\hat{F}$ .

The outline of this paper is as follows: Section 2 describes the basic assumptions and main results : the strong uniform closeness between  $F_n$  and  $\hat{F}$ , and the LIL of  $L_p$ -norm of  $\hat{F}$ . Detailed proofs are provided in Section 3.

Note that for the proof of the strong uniform closeness between  $F_n$  and  $\hat{F}$ , we use the Kiefer type approximation for the empirical process (see Csörgő and Révész (1981) [10]).

Throughout the following all limits are taken as the sample size *n* tending to  $\infty$ .

#### 2. Assumptions and the Main Results

In this section, we start with the assumptions for the kernel function *k*.

**Assumption 1.** *k*: Functions k(x), xk(x) and  $x^2k(x)$  are integrable on the whole real line and satisfy the following properties:

$$k(x) \ge 0, \ \int_{-\infty}^{+\infty} k(x)dx = 1, \\ \int_{-\infty}^{+\infty} xk(x)dx = 0 \ and \ \int_{-\infty}^{+\infty} x^2k(x)dx < \infty.$$

About the band width *h*, we assume

$$h^{3/2}\log(\log n) \to 0$$
 and  $nh^4/\log(\log n) \to 0$ , (7)

which are stronger than the assumption  $nh^4 \rightarrow 0$  used in Cheng (2017) [9].

Under the above assumptions, we first state the result for evaluating the uniform closeness between  $\hat{F}$  and  $F_n$ , which improves Theorem 2.1 in Cheng (2017) [9].

**Theorem 1.** Assume that Assumption k and (7) hold. Then, for the continuous CDF F with bounded second order derivative, we have

$$\sup_{x \in R} \sqrt{\frac{n}{\log(\log n)}} |\hat{F}(x) - F_n(x)| \to 0, \quad a.s.$$
(8)

Together with LIL in (2), the LIL can be extended to  $\hat{F}$ , as follows:

**Corollary 1.** Under the assumptions of Theorem 1, for the continuous CDF F with bounded second order derivative, we have

$$\lim \sup_{n \to \infty} \sqrt{\frac{n}{2\log(\log n)}} \sup_{x} \left| \hat{F}(x) - F(x) \right| = \frac{1}{2} \quad a.s.$$
(9)

Remark 1. Using a different approach, (9) was verified in Winter (1979) [7].

Combining (8) with (5), the LIL for  $L_p$ -norm of  $F_n$  can be extended to  $\hat{F}$ .

**Theorem 2.** Under the assumptions of Theorem 1, for any  $p \ge 1$  and the continuous CDF F with bounded second order derivative, we have

$$\lim \sup_{n \to \infty} \sqrt{\frac{n}{2\log(\log n)}} \Big[ \int_{-\infty}^{\infty} |\hat{F}(x) - F(x)|^p dF(x) \Big]^{1/p} = C(p) \quad a.s., \tag{10}$$

where C(p) is defined in (4).

**Remark 2.** Applying the facts  $C(1) = \frac{\sqrt{3}}{6}$  and  $C(2) = \frac{1}{\pi}$ , Theorem 2 can result in the following corollary:

**Corollary 2.** Under the assumptions of Theorem 1, for the continuous CDF F with bounded second order derivative, we have

$$\limsup_{n \to \infty} \sqrt{\frac{n}{2\log(\log n)}} \int_{-\infty}^{\infty} |\hat{F}_n(x) - F(x)| dF(x) = \frac{\sqrt{3}}{6} \qquad a.s.$$

and

$$\lim \sup_{n \to \infty} \sqrt{\frac{n}{2 \log(\log n)}} \Big[ \int_{-\infty}^{\infty} |\hat{F}_n(x) - F(x)|^2 dF(x) \Big]^{1/2} = \frac{1}{\pi} \qquad a.s.$$

Detailed proofs of the above results are given below.

#### 3. Proof

Set

$$U_n(x) := \frac{1}{n} \sum_{i=1}^n \{ I(X_i \le x) - F(x) \}, \quad x \in \mathbb{R}.$$

Therefore, (2) guarantees that

$$\lim \sup_{n \to \infty} \sqrt{\frac{n}{\log(\log n)}} \sup_{x} |U_n(x)| = O(1) \quad a.s.$$
(11)

For independent uniform [0, 1] random variables:  $\xi_1, \xi_2, \cdots, \xi_n$ , we define

$$V_n(v):=rac{1}{n}\sum_{i=1}^n \Big[I(\xi_i\leq v)-v\Big], \qquad orall v\in [0,1].$$

Then,  $V_n(v)$  is a standardized uniform [0,1] empirical process, and  $U_n(x)$  has the same distribution as  $V_n(F(x))$ . Using Theorem 4.4.3 and Theorem 1.15.2 in Csörgő and Révész (1981) [10], applying the Kiefer type approximation of the empirical process, there exists a Kiefer process { $K(s;t) : 0 \le s \le 1, 0 \le t < \infty$ } such that

$$\sup_{x} |nU_n(x) - K(F(x), n)| = O((\log n)^2) \quad \text{a.s.,}$$
(12)

with  $B_n(v) = K(v, n) / \sqrt{n}, 0 \le v \le 1$  being a Brownian bridge.

The Proof of Theorem 1 involves three parts: (i) applying the the triangular inequality to the distribution functions, (ii) using the the Kiefer approximation of the empirical process, and (iii) applying the the Taylor expansion. See below for details.

**Proof of Theorem 1.** Rewrite  $\hat{F}(x)$ ,

$$\hat{F}(x) = n^{-1} \sum_{i=1}^{n} \int_{-\infty}^{x} k_h (u - \varepsilon_i) du = n^{-1} \sum_{i=1}^{n} G\left(\frac{x - \varepsilon_i}{h}\right), \tag{13}$$

where  $G(x) = \int_{-\infty}^{x} k(u) du$ . By the definition of  $F_n(x)$ , performing integration by parts and a change of variable  $u = \frac{x-t}{h}$ , we can continue to rewrite  $\hat{F}(x)$ , as follows:

$$\hat{F}(x) = \int_{-\infty}^{+\infty} G\left(\frac{x-t}{h}\right) dF_n(t)$$

$$= G\left(\frac{x-t}{h}\right) F_n(t) \Big|_{-\infty}^{+\infty} + \int_{-\infty}^{+\infty} F_n(t) k\left(\frac{x-t}{h}\right) \frac{1}{h} dt$$

$$= \int_{-\infty}^{+\infty} F_n(t) \frac{1}{h} k\left(\frac{x-t}{h}\right) dt = \int_{-\infty}^{+\infty} F_n(x-hu) k(u) du.$$
(14)

Combining (14) with the properties  $k(u) \ge 0$  and  $\int_{-\infty}^{+\infty} k(u) du = 1$ , and applying the triangular inequality, we have that

$$\begin{aligned} \left| \hat{F}(x) - F_n(x) \right| &= \left| \int_{-\infty}^{+\infty} [F_n(x - hu) - F_n(x)] k(u) du \right| \\ &\leq \left| \int_{-\infty}^{+\infty} \left\{ [F_n(x - hu) - F(x - hu)] - [F_n(x) - F(x)] \right\} k(u) du \right| \\ &+ \left| \int_{-\infty}^{+\infty} [F(x - hu) - F(x)] k(u) du \right| \\ &= \left| \int_{-\infty}^{+\infty} [U_n(x - hu) - U_n(x)] k(u) du \right| + \left| \int_{-\infty}^{+\infty} [F(x - hu) - F(x)] k(u) du \right|. \end{aligned}$$

Moreover, this results in

$$\begin{split} \sup_{x} \left| \hat{F}(x) - F_{n}(x) \right| \\ &\leq \sup_{x} \left| \int_{-\infty}^{+\infty} [U_{n}(x - hu) - U_{n}(x)]k(u)du \right| + \sup_{x} \left| \int_{-\infty}^{+\infty} [F(x - hu) - F(x)]k(u)du \right| \\ &=: D_{1n} + D_{2n}, \quad \text{say.} \end{split}$$

Thus, to show Theorem 1, it is sufficient to verify that

$$\lim_{n \to \infty} \sup_{n \to \infty} \sqrt{\frac{n}{\log(\log n)}} D_{1n} = 0 \text{ and } \lim_{n \to \infty} \sqrt{\frac{n}{\log(\log n)}} D_{2n} = 0 \text{ a.s.}$$
(15)

By  $log(logn) \rightarrow \infty$  and the integrability of k(u), it follows that

$$\left(\int_{\log(\log n)}^{+\infty} + \int_{-\infty}^{-\log(\log n)}\right) k(u) du = o(1).$$
(16)

Partitioning the integral in  $D_{1n}$  into three parts, and using the triangular inequality, we can obtain

$$D_{1n} = \sup_{x} \left| \left( \int_{-\infty}^{-\log(\log n)} + \int_{\log(\log n)}^{+\infty} + \int_{-\log(\log n)}^{-\log(\log n)} \right) [U_{n}(x - hu) - U_{n}(x)]k(u)du \right|$$

$$\leq \sup_{x} \left( \int_{-\infty}^{-\log(\log n)} + \int_{\log(\log n)}^{+\infty} + \int_{-\log(\log n)}^{\log(\log n)} \right) \left| U_{n}(x - hu) - U_{n}(t) \right| k(u)du$$

$$\leq 2 \left( \sup_{x} |U_{n}(x)| \right) \left( \int_{-\infty}^{-\log(\log n)} + \int_{\log(\log n)}^{+\infty} \right) k(u)du + \sup_{x} \int_{-\log(\log n)}^{\log(\log n)} \left| U_{n}(x - hu) - U_{n}(x) \right| k(u)du$$

$$=: D_{11n} + D_{12n}, \quad \text{say.}$$
(17)

It is easy to see that (11) and (16) imply that

$$\lim \sup_{n \to \infty} \sqrt{\frac{n}{\log(\log n)}} D_{11n} = 0 \quad \text{a.s.}$$
(18)

As for  $D_{12n}$ , with the triangular inequality, (12),  $\int_{-log(logn)}^{log(logn)} k(u) du \le 1$  and the continuity of modulus of  $B_n(F(x)) = K(F(x), n)/\sqrt{n}$ , we have

$$\begin{aligned} & D_{12n} \\ \leq & \sup_{x} \frac{1}{n} \Big| \int_{-\log(\log n)}^{\log(\log n)} \left\{ [nU_{n}(x - hu) - K(F(x - hu), n)] - [nU_{n}(x) - K(F(x), n)] \right\} k(u) du \Big| \\ & + \sup_{x} \frac{1}{n} \Big| \int_{-\log(\log n)}^{\log(\log n)} [K(F(x - hu), n) - K(F(x), n)] k(u) du \Big| \\ \leq & \frac{2}{n} \Big( \sup_{x} \Big| nU_{n}(x) - K(F(x), n) \Big| \Big) + O(h \log(\log n) \sqrt{h \log(\log n)} / \sqrt{n} \\ = & O(\frac{(\log n)^{2}}{n}) + O\Big( h \log(\log n) \sqrt{h \log(\log n)} / \sqrt{n} \Big) \quad \text{a.s.} \end{aligned}$$

Hence, combining the above bound with  $\frac{(\log n)^2}{\sqrt{n \log(\log n)}} \to 0$  and the assumption  $h^{3/2} \log(\log n) \to 0$ , it follows that

$$\lim \sup_{n \to \infty} \sqrt{\frac{n}{\log(\log n)}} D_{12n} = 0. \quad \text{a.s.}$$
<sup>(19)</sup>

Next, we proceed to evaluate  $D_{2n}$ . Using the Taylor expansion with integral remainder, the properties  $\int_{-\infty}^{+\infty} uk(u)du = 0$ ,  $\int_{-\infty}^{+\infty} u^2k(u)du < \infty$ ,  $\sup_t |f'(t)| < \infty$  and  $nh^4 / \log(\log n) \to 0$ , we obtain

$$F(x - hu) - F(x) = -huf(x) + \int_{x}^{x - hu} f'(s)(x - hu - s)ds$$
  

$$= -huf(x) + \int_{0}^{-hu} f'(x - hu - t)tdt,$$
  

$$D_{2n} = \sup_{x} \left| \int_{-\infty}^{+\infty} [\int_{0}^{-hu} f'(x - hu - t)tdt]k(u)du \right|$$
  

$$\leq \sup_{x} |f'(x)| \int_{-\infty}^{+\infty} \frac{1}{2}h^{2}u^{2}k(u)du = O(h^{2}),$$
  

$$\sqrt{\frac{n}{\log(\log n)}}D_{2n} = O(\sqrt{\frac{nh^{4}}{\log(\log n)}}) \to 0.$$
(20)

Therefore, (15) can be produced from (17)–(20). We have completed the proof of Theorem 1.  $\Box$ 

**Proof of Corollary 1.** Decomposing  $\hat{F}(x) - F(x)$  into two parts and using the triangular inequality, we have

$$\begin{aligned} |\hat{F}(x) - F(x)| &= |\hat{F}(x) - F_n(x) + F_n(x) - F(x)| \\ &\leq |\hat{F}(x) - F_n(x)| + |F_n(x) - F(x)|, \\ |\hat{F}(x) - F(x)| &\geq -|\hat{F}(x) - F_n(x)| + |F_n(x) - F(x)|. \end{aligned}$$

Then, it follows that

$$\sup_{x \in R} |\hat{F}(x) - F(x)| \leq \sup_{x \in R} |\hat{F}(x) - F_n(x)| + \sup_{x \in R} |F_n(x) - F(x)|$$

and

$$\sup_{x\in R} \left| \hat{F}(x) - F(x) \right| \geq -\sup_{x\in R} \left| \hat{F}(x) - F_n(x) \right| + \sup_{x\in R} \left| F_n(x) - F(x) \right|.$$

Combining the above inequalities with (8) and (2), this guarantees that

$$\lim \sup_{n \to \infty} \sqrt{\frac{n}{2\log(\log n)}} \sup_{x} \left| \hat{F}(x) - F(x) \right| = \frac{1}{2} \quad a.s.$$

Thus, we have finished the proof of Theorem 1.  $\Box$ 

**Proof of Theorem 2.** For any  $p \ge 1$ , using the triangular inequality and the fact that  $\int_{-\infty}^{\infty} 1 dF(x) = 1$ , we have that

$$\left[\int_{-\infty}^{\infty} |\hat{F}(x) - F(x)|^{p} dF(x)\right]^{1/p} \\ \leq \left[\int_{-\infty}^{\infty} |\hat{F}(x) - F_{n}(x)|^{p} dF(x)\right]^{1/p} + \left[\int_{-\infty}^{\infty} |F_{n}(x) - F(x)|^{p} dF(x)\right]^{1/p} \\ \leq \sup_{x \in R} |\hat{F}(x) - F_{n}(x)| + \left[\int_{-\infty}^{\infty} |F_{n}(x) - F(x)|^{p} dF(x)\right]^{1/p}$$
(21)

and

$$\left[\int_{-\infty}^{\infty} |\hat{F}(x) - F(x)|^{p} dF(x)\right]^{1/p} \\ \geq -\left[\int_{-\infty}^{\infty} |\hat{F}(x) - F_{n}(x)|^{p} dF(x)\right]^{1/p} + \left[\int_{-\infty}^{\infty} |F_{n}(x) - F(x)|^{p} dF(x)\right]^{1/p} \\ \geq -\sup_{x \in \mathbb{R}} |\hat{F}(x) - F_{n}(x)| + \left[\int_{-\infty}^{\infty} |F_{n}(x) - F(x)|^{p} dF(x)\right]^{1/p}.$$
(22)

Note that we have proven

$$\sup_{x\in R}\sqrt{\frac{n}{\log(\log n)}}|\hat{F}(x)-F_n(x)|\to 0, \quad \text{ a.s}$$

in Theorem 1. Thus, combining it with (21), (22) and the law of the iterated logarithm for  $L_p$ -norm of  $F_n(x)$  in (5), it follows that

$$\lim \sup_{n \to \infty} \sqrt{\frac{n}{2\log(\log n)}} \Big[ \int_{-\infty}^{\infty} |\hat{F}(x) - F(x)|^p dF(x) \Big]^{1/p} = C(p) \ a.s.$$

Therefore, we have finished the proof of Theorem 2.  $\Box$ 

**Proof of Corollary 2.** Corollary 2 is the special case of results of Theorem 2 with p = 1, 2.  $\Box$ 

Funding: This research received no external funding.

**Data Availability Statement:** No new data were created or analyzed in this study. Data sharing is not applicable to this article.

**Acknowledgments:** The author is grateful to the Editor and two referees for their helpful comments and suggestions, which greatly improved the presentation of this article.

Conflicts of Interest: The author declares no conflict of interest.

#### References

- 1. Fabian, V.; Hannan, J. Introduction to Probability and Mathematical Statistics; Cengage Learning: Belmont, CA, USA, 1985.
- Smirnov, N.V. An approximation to the distribution laws of random quantities deter- mined by empirical data. Uspehi. Mat. Nauk. 1944, 10, 179–206.
- 3. Chung, K.L. An estimate concerning the Kolmogorov limit distribution. *Trans. Am. Math. Soc.* 1949, 67, 36–50.
- 4. Finkelstein, H. The law of the iterated logarithm for empirical distributions. Ann. Math. Statist. 1971, 42, 607–615. [CrossRef]
- 5. Gajek, L.; Kahszka, M.; Lenic, A. The law of the iterated logarithm for Lp-norms of empirical processes. *Statist. Probab. Lett.* **1996**, 28, 107–110. [CrossRef]
- 6. Yamato, H. Uniform convergence of an estimator of a distribution function. *Bull. Math. Statist.* **1973**, *15*, 69–78. [CrossRef] [PubMed]
- 7. Winter, B.B. Convergence Rate of Perturbed Empirical Distribution Functions. J. Appl. Probab. 1979, 16, 163–173. [CrossRef]
- 8. Wang, J.; Cheng, F.; Yang, L. Smooth simultaneous confidence band for cumulative distribution function. *J. Nonparametr. Stat.* **2013**, *25*, 395–407. [CrossRef]
- 9. Cheng, F. Strong uniform consistency rates of kernel estimators of cumulative distribution functions. *Commun. Stat.—Theory Methods* **2017**, *46*, 6803–6807. [CrossRef]
- 10. Csörgő, M.; Révész, P. Strong Approximation in Probability and Statistics; Academic Press: New York, NY, USA, 1981.

**Disclaimer/Publisher's Note:** The statements, opinions and data contained in all publications are solely those of the individual author(s) and contributor(s) and not of MDPI and/or the editor(s). MDPI and/or the editor(s) disclaim responsibility for any injury to people or property resulting from any ideas, methods, instructions or products referred to in the content.